Approximately optimum strata boundaries for two concomitant stratification variables under proportional allocation

Faizan Danish¹, S. E. H. Rizvi²

ABSTRACT

The proper choice of strata boundaries is an important factor determining the efficiency of the estimator of the considered characteristics of a population. In this article, the CumDepth (i=3,4) Rule for obtaining approximately optimum strata boundaries has been applied, taking into account a single-study variable along with two concomitant variables serving as the basis of the stratification variables. The relative efficiency of the proposed methods has been demonstrated theoretically and empirically by comparing them to a selection of already-existing methods in a simulation study with the use of the proportional allocation method.

Key words: stratification points, proportional allocation, minimal equation.

1. Introduction

Let there be a finite population consisting of N units, for which it is required to estimate the total or mean for the characteristic Y under study, using simple random sampling technique. In order to have this, we partition population L \times M strata:

\[ \sum_{h=1}^{L} \sum_{k=1}^{M} N_{hk} = N \]

where \( N_{hk} \) indicates the number of units in \((h,k)\)th stratum.

Let 'n' be the number of units to be drawn from the whole population and suppose that the allocation of sample size \( n_{hk} \) such that

\[ \sum_{h=1}^{L} \sum_{k=1}^{M} n_{hk} = n \]
Let $Y_{hi}$ $(i = 1, 2, 3, \ldots, N_{hi})$ be the population unit and then the population total is

$$Y = \sum_{h=1}^{L} \sum_{k=1}^{M} \sum_{i=1}^{N_{hi}} Y_{hi}$$

For the study variable, the unbiased estimate of $Y$, $\bar{Y}_{nt} = \sum_{h=1}^{L} \sum_{k=1}^{M} W_{hk} \bar{Y}_{hk}$, where $\bar{Y}_{hk} = \frac{1}{n_{hk}} \sum_{i=1}^{N_{hi}} y_{hi}$ and $W_{hk}$ denotes the weight of the $(h,k)^{th}$ stratum.

For stratified simple random sampling, the sample estimate $\bar{Y}_{nt}$ is unbiased with sampling variance as below:

$$V(\bar{Y}_{nt}) = \sum_{h} \sum_{k} (1 - f_{hk}) \frac{W_{hk}^2 \sigma_{hky}^2}{n_{hk}}$$

where $f_{hk} = \frac{n_{hk}}{N_{hk}}$ and if f.p.c is ignored, we have

$$V(\bar{Y}_{nt}) = \sum_{h} \sum_{k} W_{hk}^2 \sigma_{hky}^2 \frac{n_{hk}}{n_{hk}}$$

$\sigma_{hky}^2$ represents the population variance for the character $Y$ and is defined as

$$\sigma_{hky}^2 = \frac{1}{N_{hk}} \sum_{i=1}^{N_{hi}} (y_{hi} - \bar{Y}_{hk})^2$$

$\bar{Y}_{hk}$ being the population mean of all the $N_{hk}$ units in the $(h,k)^{th}$ stratum.

Construction of stratification points was pioneered by Dalenius (1950), while minimizing variance set of equations as the functions of population parameters were obtained and due to their implicit nature, it becomes complicated to obtain solutions. Cochran (1961, 1963) has also discussed the cases regarding the optimum boundaries. Yadav and Singh (1984), Rizvi et al. (2000), Danish et al. (2020), Khan et al. (2008), Khan et al. (2014), Danish et al. (2017), Danish et al. (2018), Danish and Rizvi (2018, 2019) and Danish, F. (2018). Rizvi and Danish (2018) made an attempt to summarise the proposed contribution towards obtaining stratification points.

The allocation procedure in which a sample size is selected as per proportion of the stratum is known as proportional allocation. In such allocation, the sample size is selected as

$$n_{hk} = \frac{n N_{hk}}{N} = n W_{hk}$$
thus, \( n_{hk} \propto N_{hk} \) and \( \sum_{h=1}^{L} \sum_{k=1}^{M} n_{hk} = n \)

Hence, under such allocation, variance is

\[
V\left(\bar{y}_{st}\right) = \frac{1}{n} \sum_{h} \sum_{k} W_{hk} \sigma_{hky}^2
\]

In this paper, for obtaining stratification points using classical approach for two concomitant variables as the basis of stratification variables and a single study variable under the proportional allocation method by assuming different distributions of the concomitant variables and both dependent and independent cases have been discussed as well.

2. Variance expression

Let the regression model of the response variable Y and the two information variables X & Z be given as

\[ Y = C(X, Z) + e \]

where 'e' is error term such that

\[ E(e|X, Z) = 0 \quad \text{and} \quad V(e|X, Z) = \eta(X, Z) > 0, \forall \ x \in (a, b), \ z \in (c, d), \]

\[ (b - a) < \infty, (c - d) < \infty \]

If joint marginal of X and Z is \( f(x, z) \) and \( f(x) \) and \( f(z) \) denotes marginal densities of individual variables, respectively, then under above regression model,

we have \( W_{hk} = \int_{x_{h-1}}^{x_h} \int_{z_{k-1}}^{z_k} f(x, z) \partial x \partial z \) is weight of a stratum.

\[ \mu_{hky} = \mu_{hkc} = \frac{1}{W_{hk}} \int_{x_{h-1}}^{x_h} \int_{z_{k-1}}^{z_k} c(x, z) f(x, z) \partial x \partial z \]

\[ \sigma^2_{hky} = \sigma^2_{hkc} + \mu_{hkc} \]

\( \mu_{hkc} \) and \( \sigma^2_{hkc} \) denotes its mean and variation respectively,

where \( (x_{h-1}, x_h, z_{k-1}, z_k) \) be the stratification points and \( \mu_{hkc} \) is the average value of the function \( \eta(x, z) \) and \( \sigma^2_{hkc} \) as

\[
\sigma^2_{hkc} = \frac{1}{W_{hk}} \int_{x_{h-1}}^{x_h} \int_{z_{k-1}}^{z_k} c^2(x, z) f(x, z) \partial x \partial z - (\mu_{hkc})^2
\]
Using these relations, the variance can be expressed in terms of the population parameters of the function of X and Y and \( V(e|z) \). The variance expression for the case of proportional allocation is therefore given by

\[
V(\bar{y}_{st})_{\text{prop}} = \frac{\sum \sum W_{hk} \left( \sigma^2_{hkc} + \mu_{hk}\eta \right)}{n}
\]  

(2.1)

The expression for various terms can be in terms of Singh and Sukhatme (1969) and Danish et al. (2018).

3. Minimal equations for proportional allocation

Since \( \sum \sum W_{hk}\mu_{hk}\eta = \mu\eta \), which is the population parameter and therefore is a fixed constant. Hence, minimization of (2.1) is equivalent to minimization

\[
V_p = \sum \sum W_{hk} \sigma^2_{hkc}
\]

(3.1)

Thus, to obtain minimal equations, we minimize \( V_p \) by with respect of \( x_h \) and equate to zero, we get

\[
\frac{\partial}{\partial x_h} V_p = \sum_k \left[ W_{hk} \frac{\partial}{\partial x_h} \sigma^2_{hkc} + \sigma^2_{hkc} \frac{\partial}{\partial x_h} W_{hk} + W_{ik} \frac{\partial}{\partial x_h} \sigma^2_{ikc} + \sigma^2_{ikc} \frac{\partial}{\partial x_h} W_{ik} \right] = 0
\]

After further simplification, we get

\[
\sum_k \left[ W_{hk} f_{x_k}^{x_h} \frac{f(x_h, z)}{W_{hk}} \left( c(x_h, z) - \mu_{hkc} \right)^2 - \sigma^2_{hkc} \right] \frac{\partial}{\partial x_h} + \sigma^2_{hkc} \sum_k f(x_h, z) \right] = 0
\]

(3.2)

For obtaining minimal equations we also differentiate \( V_p \) partially w.r.t. \( z_k \) in a similar way, we get

\[
\sum_h \left[ W_{hk} f_{z_k}^{x_h} \frac{f(x, z_k)}{W_{hk}} \left( c(x, z_k) - \mu_{hkc} \right)^2 - \sigma^2_{hkc} \right] \frac{\partial}{\partial z_k} + \sigma^2_{hkc} \sum_h f(x, z_k) \right] = 0
\]

(3.3)
However, for obtaining minimal equations we minimize $V_p$ on equating the partial derivative of this expression with respect of $x_h$ and $z_k$ to zero, we get

$$W_{hk} f(x_h, z_k) \left\{ \frac{[c(x_h, z_k) - \mu_{hkc}]^2 - \sigma_{hkc}^2}{W_{hk}} + f(x_h, z_k)\sigma_{hkc}^2 \right\} = W_{ij} f(x_h, z_k) \left\{ \frac{[c(x_h, z_k) - \mu_{ijc}]^2 - \sigma_{ijc}^2}{W_{ij}} + f(x_h, z_k)\sigma_{ijc}^2 \right\}$$

This gives the equation as

$$c(x_h, z_k) = \frac{(\mu_{hkc} + \mu_{ijc})^2}{2}, \quad i = h + 1, h = 1, 2, ..., L - 1 \quad j = k + 1, k = 1, 2, ..., M - 1$$

(3.4)

On the condition that $\lambda(x, z) = c'(x, z)f(x, z)$ belongs to class $\Omega$ functions, solutions to the system of equation (3.4) give OSB in the sense of minimization of variance $V(\tilde{\nu}_{st})_{prop}$. These equations are also very difficult to solve and, therefore, for these equations also we shall find methods of obtaining approximation to the exact solutions $[x_h, z_k]$. Further better approximation can be obtained by using some approximate iterative procedures.

4. Some miscellaneous results

In the case of complexities in the equations, let us impose few regularity conditions on $f(x, z), c(x, z)$ and $\eta(x, z)$. We state that $\zeta(x, z)$ belongs to class $\Omega$ if it satisfies

i) $0 < \zeta(x, z)$

ii) $\zeta(x, z) < \infty$

iii) $\zeta(x, z), \zeta'(x, z)$ and $\zeta''(x, z)$ exist and are continuous $\forall (x, z)$ in $[(a, b), (c, d)]$ respectively such that $(b - a) < \infty$ and $(d - c) < \infty$.

Let us suppose $f(x, z)$ and $\eta(x, z)$ belong to class $\Omega$ and the function $c(x, z)$ satisfies the conditions (ii) and (iii).
Before we proceed to prove the results, let us define the symbol 'O', which has been
used in the present investigation.

For two functions \( T_1(x, z) \) and \( T_2(x, z) \), such that the ratio \( T_1(x, z)/T_2(x, z) \) remains bounded as \( x \) and \( z \) tends to their limits, then we can write \( T_1(x, z) = O(T_2(x, z)) \).

**Lemma 4.1:** If the function \( I_{ij}(x, z) \) is defined as

\[
I_{ij}(x, z) = \int_{x_1}^{x_2} \int_{z_1}^{z_2} \left( t_1 - x_1 \right) ^i \left( t_2 - z_1 \right) ^j f(t_1, t_2) \partial t_1 \partial t_2 , \quad x_1 < x_2 \text{ and } z_1 < z_2
\]

then

\[
I_{ij}(x, z) = \left[ k_1^{i+1}k_2^{j+1} \over (i+1)(j+1) f + k_1^{i+1}k_2^{j+2} \over (i+1)(j+2) f_x + k_1^{i+2}k_2^{j+1} \over (i+2)(j+1) f_z \right.
\]

\[
+ \frac{1}{2!} \left[ k_1^{i+3}k_2^{j+2} \over (i+3)(j+2) f_{xx} + k_1^{i+2}k_2^{j+3} \over (i+2)(j+3) f_{zz} \right] + O(k^{i+j+5})
\]

(4.1)

where \( f(t_1, t_2) = f, \frac{\partial f}{\partial t_1} = f_x, \frac{\partial f}{\partial t_2} = f_z, \frac{\partial^2 f}{\partial t_1^2} = f_{xx}, \frac{\partial^2 f}{\partial t_2^2} = f_{zz}, \frac{\partial^2 f}{\partial t_1 \partial t_2} = f_{xz} \),

\( k_1 = x_2 - x_1 \) and \( k_2 = z_2 - z_1 \).

**Lemma 4.2:** Let \( \mu_\eta(x, z) \) denote the conditional expectation of the function \( \eta(t_1, t_2) \), so that

\[
\mu_\eta(x, z) = \frac{\int_{x_1}^{x_2} \int_{z_1}^{z_2} \eta(t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2}{\int_{x_1}^{x_2} \int_{z_1}^{z_2} f(t_1, t_2) \partial t_1 \partial t_2}
\]

Then, the series expansion of \( \mu_\eta(x, z) \) at point \( (t_1, t_2) \) is given by

\[
\mu_\eta(x, z) = \eta + \left[ \frac{\eta'(f_{xx} + f_{xz}) + 2 f \eta'}{12 f \eta} \right] (k_1 + k_2)^2
\]

\[
+ O((k_1 + k_2)^3)
\]

(4.2)
Lemma 4.3: If \( \sigma^2_\eta(x, z) \) denotes the conditional variance of the function \( \eta(t_1, t_2) \) defined in the interval \((x, z)\), then
\[
\sigma^2_\eta(x, z) = \frac{(k)^2}{12} \left( \frac{\eta}{\eta^*} \right)^2 \left[ 1 + \frac{\eta^*}{\eta} (k)^1 + O(k)^2 \right] \tag{4.3}
\]
where \((k)^1\) and \((k)^2\) denote all \(k\)'s with power '1' and '2' respectively.

Lemma 4.4:
\[
(k_1 k_2)^{\frac{1}{2}} \int_{z_1}^{z_2} \int_{x_1}^{x_2} f(t_1, t_2) \frac{\partial t_1 \partial t_2}{\partial t_1 \partial t_2} = \left[ \int_{z_1}^{z_2} \int_{x_1}^{x_2} \sqrt{f(t_1, t_2) \partial t_1 \partial t_2} \right]^{\frac{1}{2}} \left[ 1 + O(k^2) \right] \tag{4.4}
\]

Lemma 4.5: With \( I_{00}(x, z) \) and \( \sigma^2_\eta(x, z) \) defined as in Lemma 4.1 and Lemma 4.3 respectively, we have
\[
I_{00}(x, z) \sigma^2_\eta(x, z) = \frac{k^2}{12} \int_{z_1}^{z_2} \int_{x_1}^{x_2} \eta^2(t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2 \tag{4.5}
\]
where \(k\) denotes any \(k_h\) or \(k_k\).

5. Minimal equations and their approximate solutions

In this section we will obtain expansion of the series given in (3.4) about the points \(x_h\) and \(z_k\) the common boundary of \((h, k)^{th}\) and \((h+1, k+1)^{th}\) strata and obtain the approximate systems of equation, which will give approximately optimum points of stratification as their solutions. In doing so we shall make use of the Lemma’s already 4.1-4.5.

The minimal equations for this method are given by
\[
c(x_h, z_k) - \mu_{hkc} = \mu_{ijc} - c(x_h, z_k)
\]
\[
i = h+1, h = 1, 2, \ldots L, j = k+1, k = 1, 2, \ldots, M
\]
For R.H.S., the corresponding expansion for the L.H.S. may be obtained by changing the signs of coefficients of even powers of \((k_i, k_j)\) the width of \((i, j)^{th}\) stratum, where \(k_i = x_{h+1} - x_h\), \(k_j = z_{k+1} - z_k\). From (4.2), after replacing \((x_1, x_2)\) by \((x_h, x_{h+1})\) and \((z_1, z_2)\) by \((z_k, z_{k+1})\), we have

\[
\mu_{k'} = c \left[ 1 + \frac{c'}{2c} k_i + \left( \frac{c' f_x + 2 f c^*}{12 f c} \right) k_i^2 + \left( \frac{f f_x c' + f f_x c + f^2 c^*}{24 f^2 c} \right) k_i^3 + O(k_i^4) \right]
\]

where \(k_i = x_{h+1} - x_h\) and derivatives are evaluated at \(x_h\).

However, when the same functions are differentiated w.r.t. \(z_k\), we have

\[
\mu_{k'} = c \left[ 1 + \frac{c'}{2c} k_j + \left( \frac{c' f_z + 2 f c^*}{12 f c} \right) k_j^2 + \left( \frac{f f_z c' + f f_z c + f^2 c^*}{24 f^2 c} \right) k_j^3 + O(k_j^4) \right]
\]

where \(k_j = z_{k+1} - z_k\).

Here, the derivatives are evaluated at both \(x_h\) and \(z_k\), we get

\[
\mu_{k'} = c \left[ 1 + \frac{c'}{2c} (k_i k_j) + \left( \frac{c' (f_x + f_z) + 2 f c^*}{12 f c} \right) (k_i k_j)^2 + (f_4)(k_i k_j)^3 + O(k_i k_j)^4 \right]
\]

\[
f_4 = f \left(f_{xx} + f_{zz} + f_{xz}\right) c' + f \left(f_x + f_z\right) c^* + f^2 c^*
\]

where

Thus, we have

\[
\mu_{k'} = c \left[ 1 + \frac{c'}{2c} (k_i k_j) + \left( \frac{c' (f_x + f_z) + 2 f c^*}{12 f} \right) (k_i k_j)^2 + O(k_i k_j)^2 \right]
\]

Similarly, we get

\[
\mu_{k'} = c \left[ 1 + \frac{c'}{2c} (k_i k_j) + \left( \frac{c' (f_x + f_z) + 2 f c^*}{12 f} \right) (k_i k_j)^2 + O(k_i k_j)^2 \right]
\]
Evaluating derivatives at $x_h$ and $z_k$. Therefore, equation (3.4) can be put as

$$
\frac{(k_hk_k}{2} c + \left(\frac{c'(f_x + f_z) + 2fc^*}{12f}\right)(k_hk_k) + O(\epsilon^2) + O(\epsilon^2) = \frac{(k_hk_k}{2} c + \left(\frac{c'(f_x + f_z) + 2fc^*}{12f}\right)(k_hk_k) + O(\epsilon^2)
$$

(5.1)

Now, let us consider an expansion of the function

$$
B_{hk} = \int_{z_{k-1}}^{z_k} \int_{x_{h-1}}^{x_h} c' (t_1, t_2) f(t_1, t_2) \, dt_1 \, dt_2
$$

about the point $[x_h, z_k]$. Expanding the integral about $x_h$ and $z_k$ with the help of the Taylor’s expansion for two variables, we have

$$
B_{hk} = fc' k_hk_k k \left[1 - \frac{(f_x + f_z)c' + 2fc^*}{2fc'} \left(k_hk_k\right) + O(\epsilon^2)\right]
$$

(5.2)

where in (5.2) also the function of $f, \eta$ and their derivatives are evaluated at $x_h$ and $z_k$. Thus, we find that

$$
\frac{(k_hk_k}{8} B_{hk} c' = \frac{(k_hk_k}{8} \left[1 - \frac{(f_x + f_z)c' + 2fc^*}{2fc'} \left(k_hk_k\right) + O(\epsilon^2)\right]
$$

or

$$
\left[\frac{(k_hk_k}{8} B_{hk} c'\right]^1 = \left[\frac{(k_hk_k}{8} \left[1 - \frac{(f_x + f_z)c' + 2fc^*}{2fc'} \left(k_hk_k\right) + O(\epsilon^2)\right]
$$

(5.3)

Similarly, we obtain

$$
\left[\frac{(k_hk_k}{8} B_{k} c'\right]^1 = \left[\frac{(k_hk_k}{8} \left[1 - \frac{(f_x + f_z)c' + 2fc^*}{6fc'} \left(k_hk_k\right) + O(\epsilon^2)\right]
$$
Therefore, the minimal equations (5.1) can be put as

\[
\left[ \frac{(k_hk_k)^2}{8} \right] \frac{1}{3} + O(k_hk_k)^2 = \left[ \frac{(k_jk_j)^2}{8} \right] \frac{1}{3} + O(k_jk_j)^2
\]

(5.4)

Hence, if the terms of order \( \text{Sup} \left( ((a,b),(c,d)) (k_hk_j) \right)^3 \) can be neglected, we can replace the minimal equations approximately by

\[
\left[ \frac{(k_hk_k)^2}{8} \frac{B_{hk}c}{f} \right] \frac{1}{3} = \left[ \frac{(k_jk_j)^2}{8} \frac{B_{j}c}{f} \right] \frac{1}{3}
\]

or \( (k_hk_k)^2 B_{hk} = \text{Constant} \) (5.5)

In the case when it is possible to find a function \( Q(x_{h-1}, x_h, z_{k-1}, z_k) \) such that

\[
(k_hk_k)^2 B_{hk} = (k_hk_k)^2 \int_{z_{k-1}}^{z_k} \int_{x_{h-1}}^{x_h} c^2 (t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2 = Q(x_{h-1}, x_h, z_{k-1}, z_k) \left[ 1 + O(k_hk_k)^2 \right]
\]

(5.6)

Thus, the system of equations (5.5) to the same degree of accuracy can be put as

\[
Q'(x_{h-1}, x_h, z_{k-1}, z_k) = \text{Constant}
\]

(5.7)

The above results can be put in the form of a note as follows.

**Remark 1:** If the regression of the dependent variable \( Y \) and stratification variables \( X \) and \( Z \) in an unbounded population is given by

\[
Y = C(X, Z) + e
\]

where \( e \) is the error component such that \( E(e | x, z) = 0 \) and \( \text{V}(e | x, z) = \eta(x, z) > 0, \forall x \in (a, b) \) and \( z \in (c, d) \) with finite deviation of the intervals, and in addition if \( c^2 (x, z) f(x, z) \) belong to \( \Omega \), then the system of
equations (3.4) giving strata boundaries \([x_h, z_k]\), which correspond to the minimum of \(V(y_{st})_{prop}\), can be put as

\[
\left(\int_{\frac{x_k}{x_{h-1}}}^{x_h} \int_{x_{h-1}}^{x_k} \frac{c^2}{(t_1, t_2)} f(t_1, t_2) \partial t_1 \partial t_2 \left[1 + O(k_h k_k)^2\right]\right)^{\frac{1}{3}} = \left(\int_{\frac{x_k}{x_{h-1}}}^{x_h} \int_{x_{h-1}}^{x_k} \frac{c^2}{(t_1, t_2)} f(t_1, t_2) \partial t_1 \partial t_2 \left[1 + O(k_i k_j)^2\right]\right)^{\frac{1}{3}}.
\]

If the terms of order \(O\left(\frac{c}{((a, b), (c, d))}\right)\) can be neglected, these equations can be replaced by the approximate system of equations

\[
\left(\int_{\frac{x_k}{x_{h-1}}}^{x_h} \int_{x_{h-1}}^{x_k} c^2 (t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2 \left[1 + O(k_h k_k)^2\right]\right) = \text{Constant}
\]

Or equivalently by

\[
Q_1 (x_{h-1}, x_h, z_{k-1}, z_k) = \text{Constant}
\]

Therefore,

\[
Q_1 (x_{h-1}, x_h, z_{k-1}, z_k) \left[1 + O(k_h k_k)^2\right] = \left(k_h k_k\right)^2 \left[1 + O(k_h k_k)^2\right] + \text{Constant}
\]

The same result can also be obtained by minimizing the function

\[
\sum_{h, k} \int_{\frac{x_k}{x_{h-1}}}^{x_h} \int_{x_{h-1}}^{x_k} c^2 (t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2 \left[1 + O(k_h k_k)^2\right]
\]

as in the light of Lemma 5, \(12 \sum_{h, k} W_{h,k} \sigma_{h,k}^2\) equals to this function.

Thus, we find that if the function \(c^2 (x, z) f(t_1, t_2)\) belongs to the class \(\Omega\) the minimum value of \(\sum_{h} \sum_{k} W_{h,k} \sigma_{h,k}^2\) and therefore \(V(y_{st})_{prop}\) exits and the set of strata boundaries \([x_h, z_k]\), corresponding to this minimum, is the solution of the systems of equations (3.4) or equivalently of (5.4). These equations are very difficult to solve exactly and it becomes essential to find some approximation to stratification points.
It may be precisely solved by substituting the exact minimal equations by other systems of equations which are comparatively easy to solve but are only asymptotically equivalent to the exact equations. The error is introduced because we neglect terms of higher powers of the strata widths which can be justified when the total stratum is large. The approximate systems of equations are obtained by neglecting terms of order \( O\left( m^3 \right) \) where \( m = \frac{\text{Sup}}{((a,b),(c,d))} \left(k_h k_k \right)\), on both sides of (5.4). For large strata, the terms of order \( O\left( m^3 \right) \) are small and therefore the error involved in the approximate systems of equations is small, although this error is comparatively larger than the one involved in the case of optimum allocation. Here, we shall develop the approximate systems of equations given in (5.5) and (5.7).

6. Approximate systems of equations

I. If in the expansion of the minimal equations (3.4) we neglect all terms except the first on both sides of the equation, the solution is obtained by taking

\[ x_h = \text{constant} = \frac{b-a}{L}, \quad h=1,2,\ldots,L \text{ and } z_k = \text{constant} = \frac{d-c}{M}, \quad k=1,2,\ldots,M. \]

\[ (6.1) \]

Therefore

\[ x_h = a + \left(\frac{b-a}{L}\right) h \quad \text{with} \quad x_0 = a \quad \text{and} \quad x_L = b \]

and

\[ z_k = c + \left(\frac{d-c}{M}\right) k \quad \text{with} \quad z_0 = c \quad \text{and} \quad z_M = d \]

It cannot be suspected that this set of approximations can give good solutions as these are simple to obtain. However, the method is not applicable in the case of infinite range.

II. An approximation to the optimum points of stratification is obtained by solving the systems of equations

\[ \left(k_h k_k \right) \int_{x_{h-1}}^{x_h} \int_{z_{k-1}}^{z_k} c^2 \left(t_1, t_2\right) f \left(t_1, t_2\right) e^{t_1} e^{t_2} = C_1 \]

\[ (6.2) \]

as shown in (5.5). The solutions of this system of equations and also of those that will now follow, are expected to be closer to the optimum stratification points as compared to the solutions obtained from (6.1).
III. From Lemma 4 and equation (6.2), we get a general class of approximate systems of equations as

\[
\left( k_h k_k \right)^{3\lambda - 1} \int_{x_{k-1}}^{x_k} \int_{z_{k-1}}^{z_k} \left( c^2 (t_1, t_2) f(t_1, t_2) \right)^{\lambda} \partial t_1 \partial t_2 \left[ \int_{x_{k-1}}^{x_k} \int_{z_{k-1}}^{z_k} \right]^{1/\lambda} = \text{Constant}
\]

However, for \( \lambda = \frac{1}{2} \), we have

\[
\left( k_h k_k \right) \int_{x_{k-1}}^{x_k} \int_{z_{k-1}}^{z_k} c^2 (t_1, t_2) \sqrt{f(t_1, t_2)} \partial t_1 \partial t_2 \bigg]^2 = C_2
\]

and for \( \lambda = \frac{1}{3} \), we have a system of equations as

\[
\left[ \int_{x_{k-1}}^{x_k} \int_{z_{k-1}}^{z_k} 3c^2 (t_1, t_2) f(t_1, t_2) \partial t_1 \partial t_2 \bigg]^3 = C_3
\]

giving approximations to stratification points \([x_h, z_k]\). As remarked in the case of the optimum allocation method, in some particular cases some of the approximate systems given in the above equations may be meaningless. Therefore, depending upon the situation, one should make the approximate choice of the systems of equations for obtaining the approximations to optimum points \([x_h, z_k]\).

7. **Cum \(3D_3(x, z)\) Rule**

If the function \( D_3(x, z) = c^2 (x, z) f(x, z) \) is bounded and its first two derivatives exist \( \forall x \in [a, b] \) & \( z \in [c, d] \), then taking equal intervals with a given values of L and M on the cumulative cube root of \( D_3(x, z) \) will give AOSB \([x_h, z_k]\).

**Remarks:**

I. If we take either \( c(x, z) = \alpha + \beta x \) or \( c(x, z) = \alpha + \gamma z \) in \( D_3(x, z) \) it reduces to the method proposed by Singh and Sukhatme (1969).

II. If the function \( c^2 (x, z) \) is constant, therefore the proposed method reduces to Cum \(3f(x, z)\) rule.
Further, for any distribution and given number of strata the set of AOSB will remain unchanged with respect to the form of conditional variance. However, the efficiency of the stratification will differ from stratified simple random sampling estimators as well as other estimators with the choice of various forms of conditional variance.

8. Empirical study

For the purpose of empirical study, the effectiveness of the methods of finding approximation to the optimum points of stratification, we have considered the system of minimal equations obtained for the case of proportional allocation. In this illustration we shall consider equal interval approximation and the system of approximations given in (6) article. The former approximation is specially considered due to its simplicity. From all the later approximations we have only chosen one suitable method. Since the order of approximation involved in all these methods is the same, this one approximation will give the idea about the effectiveness of all other approximations given in article (6). For the sake of simplicity, the linear regression line \( Y \) on \( X \) and \( Z \) have been taken as the form \( y = \alpha + \beta x + \gamma z + e \). Here, it is considered that the two auxiliary variables used for stratification are dependent. From all the subsequent approximations we have only chosen one suitable method. Since the order of approximation involved in all these methods is the same, this one approximation will give the idea about the effectiveness of all other approximations. For obtaining the stratification points under proportional allocation let us assume \( c(x, z) = \alpha + \beta x + \gamma z \). Further, let us assume that the correlation coefficient between \( X \) and \( Z \) is denoted by \( \rho \) and is equal to 0.65. Let us consider the following examples:

**Empirical study 1:**

Suppose

\[
 f(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, \quad x \geq 0
\]

and the variable \( Z \) has

\[
 f(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad z \geq 0
\]

In order to obtain the OSB when both the variables are standard normally distributed by assuming the value of regression coefficients \( \beta = 0.65 \) and \( \gamma = 0.57 \). For obtaining total 16 strata, 4 along \( X \) variable and 4 along \( Z \) variable using the
proposed \( \sqrt{3D_3(x, z)} \) rule, by solving it in Mathematica Software assuming that the distribution of \( X \) and \( Z \) are truncated at \( x = 6 \) and \( z = 4 \), respectively, we get the stratification points as below:

Table 1. OSB and Variance, for standard normally distributed auxiliary variables

<table>
<thead>
<tr>
<th>OSB ((x_k, z_k))</th>
<th>(\sqrt{3D_3(x, z)}) Rule Variance</th>
<th>(\sqrt{3D_3(x, z)}) Rule Variance (Singh 1975)</th>
<th>% R.E.</th>
</tr>
</thead>
<tbody>
<tr>
<td>(0.3347, 0.2673)</td>
<td>0.06798628</td>
<td>0.182346</td>
<td>268.21</td>
</tr>
<tr>
<td>(0.5779, 0.2673)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1.9004, 0.2673)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(6.0000, 0.2673)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.3347, 0.5284)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.5779, 0.5284)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1.9004, 0.5284)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(6.0000, 0.5284)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.3347, 0.9865)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.5779, 0.9865)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1.9004, 0.9865)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(6.0000, 0.9865)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.3347, 4.0000)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(0.5779, 4.0000)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(1.9004, 4.0000)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>(6.0000, 4.0000)</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Empirical study 2:

Let

\[
f(x) = \begin{cases} 
\frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{\text{log}x - \mu}{2\sigma^2}}; x > 0, \sigma > 0 \\
0, else\text{where}
\end{cases}
\]

and

\[
f(z) = \begin{cases} 
\frac{1}{b - a}; a \leq z \leq b \\
0, else\text{where}
\end{cases}
\]

To obtain the OSB in proportional allocation using the proposed \( \sqrt{3D_3(x, z)} \) rule for uncorrelated auxiliary variables having densities as defined above. Standardised log-normal distribution is defined in the interval \( x \in [0,10] \) and the other variable \( z \in [0,1] \), and \( \beta=0.82 \) \& \( \gamma=0.437 \). For \( 3 \times 2 \) \((L \times M) = 6\) strata, i.e. 3 along X variable and
2 along Z variable, the results obtained after solving the function using Mathematica Software are presented in the following tables as:

<table>
<thead>
<tr>
<th>OSB</th>
<th>Variance</th>
</tr>
</thead>
<tbody>
<tr>
<td>( \left( x_h, z_k \right) )</td>
<td>( \left( \frac{3}{2} D_2(x, z) \text{ Rule} \right) )</td>
</tr>
<tr>
<td>(3.4216,0.4759)</td>
<td>0.035281796</td>
</tr>
<tr>
<td>(6.5319,0.4759)</td>
<td></td>
</tr>
<tr>
<td>(10.0000,0.4759)</td>
<td></td>
</tr>
<tr>
<td>(3.4216,1.0000)</td>
<td></td>
</tr>
<tr>
<td>(6.5319,1.0000)</td>
<td></td>
</tr>
<tr>
<td>(10.0000,1.0000)</td>
<td></td>
</tr>
</tbody>
</table>

9. For independent auxiliary variables under proportional allocation \( (\rho = 0) \)

In order to propose a technique under proportional allocation when the two auxiliary variables are independent to each other we need to proceed in the same way as proceeded in the case when they were dependent only with the difference that here in this case we have to take marginal densities rather than joint densities under consideration. We can write (5.3) as

\[
\frac{(k_h)^2}{8} \left( \frac{B_h c'(x)}{f(x)} \right)^{\frac{1}{3}} \frac{(k_h)^2}{8} \left( \frac{B_h c'(z)}{f(z)} \right)^{\frac{1}{3}}
\]

\[
= \frac{(k_h)c'(x)}{2} \left[ 1 - \left( \frac{f_x c' + 2f(x)c''(x)}{6f(x)c'(x)} \right) (k_h) + O(k_h)^2 \right] \]

\[
\frac{(k_k)c'(z)}{2} \left[ 1 - \left( \frac{f_z c' + 2f(z)c''(z)}{6f(z)c'(z)} \right) (k_k) + O(k_k)^2 \right]
\]

In a similar way, we have

\[
\frac{(k_i)^2}{8} \left( \frac{B_i c'(x)}{f(x)} \right)^{\frac{1}{3}} \frac{(k_j)^2}{8} \left( \frac{B_j c'(z)}{f(z)} \right)^{\frac{1}{3}}
\]

\[
= \frac{(k_i)c'(x)}{2} \left[ 1 - \left( \frac{f_x c' + 2f(x)c''(x)}{6f(x)c'(x)} \right) (k_i) + O(k_i)^2 \right] \]

\[
\frac{(k_j)c'(z)}{2} \left[ 1 - \left( \frac{f_z c' + 2f(z)c''(z)}{6f(z)c'(z)} \right) (k_j) + O(k_j)^2 \right]
\]
where \( B_h = \int_{x_{h-1}}^{x_h} c^2 (t_1) f(t_1) dt_1 \) and \( B_k = \int_{z_{k-1}}^{z_k} c^2 (t_2) f(t_2) dt_2 \)

However, if the terms of order \( O\left(\sup_{a,b} (k_h)\right)^3 \) and \( O\left(\sup_{c,d} (k_k)\right)^3 \) can be neglected, we replace the minimal equations approximately by

\[
\begin{align*}
\left(\frac{k_h}{8}B_{hc}(x)\right)^{\frac{1}{3}} = \frac{1}{3} & \left(\frac{k_h}{8}B_{hc}(z)\right)^{\frac{1}{3}} = \left[ \left(\frac{k_h}{8}B_{hc}(x)\right)^{\frac{1}{3}} \right]^{\frac{1}{3}} \\
\left(\frac{k_k}{8}B_{kc}(x)\right)^{\frac{1}{3}} = \frac{1}{3} & \left(\frac{k_k}{8}B_{kc}(z)\right)^{\frac{1}{3}} = \left[ \left(\frac{k_k}{8}B_{kc}(z)\right)^{\frac{1}{3}} \right]^{\frac{1}{3}}
\end{align*}
\]

or in other words we have that \( k_h^2 B_h \) and \( k_k^2 B_k \) are constants. In the case when it is possible to find a function \( \hat{Q}_1(x_{h-1}, x_h) \) and \( \hat{Q}_1(z_{k-1}, z_k) \) such that

\[
\left(\frac{k_h}{8}B_{hc}(x)\right)^{\frac{1}{3}} = \hat{Q}_1(x_{h-1}, x_h)^{\frac{1}{3}}
\]

and

\[
\left(\frac{k_k}{8}B_{hc}(z)\right)^{\frac{1}{3}} = \hat{Q}_1(z_{k-1}, x_k)^{\frac{1}{3}}
\]

The Remark 1 can be proceeded in the case of independent variables too. Similarly, an approximate system of equations can be proposed in the same way as proposed in the case when auxiliary variables are dependent, and can be written as

\[
(k_h k_k)^2 \int_{x_{h-1}}^{x_h} \int_{z_{k-1}}^{z_k} c^2(t_1) c^2(t_2) f(t_1) f(t_2) \partial t_1 \partial t_2 = C_1.
\]

The solution of this system of equations and also those that will now follow are expected to be closer to the optimum points of stratification as compared to the strata obtained composed from \( k_h = \frac{h-a}{L} = \text{Constant} \) and \( k_k = \frac{d-c}{M} = \text{Constant} \). Therefore, depending on the situation one should make the approximate choice of the system of equations for obtaining the approximation of optimum points of stratification.
10. Cum $\sqrt[3]{D_4(x,z)}$ Rule

For equal intervals with given values of L and M on the cumulative cube root of $D_4(x,z)$ will give AOSB if the function $D_4(x,z) = c^2(x)c^2(z)f(x)f(z)$ is bounded and its first derivative exists in all $x \in [a,b]$ and $z \in [c,d]$.

Remarks:
1. If the functions $c^2(x)$ and $c^2(z)$ are constants, then the proposed method is reduced to cum $\sqrt[3]{f(x)f(z)}$ rule.
2. If we take $c(x) = c(z)$ and $f(x) = f(z)$, then the proposed method is reduced to the Yadava and Singh (1984) method of Cum $\sqrt[3]{B_2(x)}$

$$B_2 = \frac{f(x)x^2c^2(x) + x\phi'(x) - \phi(x)}{x^3}$$

11. Empirical study

We shall demonstrate empirically the efficiency of the given method obtaining approximately optimum strata boundaries (AOSB). For this purpose, we have considered the system of minimal equations obtained for the case of proportional allocation when the two auxiliary variables used for stratification are independent. From all the subsequent approximations we have only chosen one suitable method. Since the order of approximation involved in all these methods is the same, this one approximation will give the idea about the effectiveness of all other approximations. For obtaining the stratification points under proportional allocation let us assume $c(x,z) = \alpha + \beta x + \gamma z$. Let us consider the following examples:

Empirical study 3:

Let $f(x) = 2(2-x), 1 \leq x \leq 2$ and $f(z) = e^{-z+1}, 1 \leq z \leq 6$

In order to obtain stratification points when the auxiliary variable X follows right-triangular distribution defined in [1,2] and auxiliary variable Z follows exponential distribution defined in [1,6] we assume the values of $\beta = 0.567$ and $\gamma = 0.257$. While execution for obtaining OSB using Cum $\sqrt[3]{D_4(x,z)}$ Rule by solving the function using Mathematica Software for 6 strata, 2 along X variable and 3 along Z variable. The results obtained are presented in the following table.
Table 3. OSB and Variance, with right-triangular and exponential distribution

<table>
<thead>
<tr>
<th>OSB ((x_h, z_k))</th>
<th>Variance</th>
<th>% R.E.</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Cum (\frac{3}{4}D_4(x, z)) Rule</td>
<td>Yadava and Singh, 1984</td>
</tr>
<tr>
<td>(1.5000, 1.9474)</td>
<td>0.089542</td>
<td>169.89</td>
</tr>
<tr>
<td>(1.0000, 1.9474)</td>
<td>0.152122</td>
<td></td>
</tr>
<tr>
<td>(1.5000, 3.3368)</td>
<td>(1.0000, 3.3368)</td>
<td></td>
</tr>
<tr>
<td>(1.5000, 6.0000)</td>
<td>(1.0000, 6.0000)</td>
<td></td>
</tr>
</tbody>
</table>

**Empirical study 4:** Let us consider the distribution of X as right-triangular having

\[ f(x) = 2(2 - x), 1 \leq x \leq 4 \]

and Z variable is having a uniformly distributed having

\[ f(z) = \frac{1}{b-a}, 1 \leq z \leq 2 \]

In order to find the OSB when one of the auxiliary variable is following right-triangular distribution and the other uniform distribution, we assume the value of \(\beta = 0.56\) and \(\gamma = 0.762\). The stratification points obtained for total 6 strata among that 3 along X variable and 2 along Z variable for the Cum \(\frac{3}{4}D_4(x, z)\) Rule using Mathematica Software for solving the function are presented in the following table.

Table 4. Uncorrelated variables having right-triangular and exponential distribution, OSB and Variance

<table>
<thead>
<tr>
<th>OSB ((x_h, z_k))</th>
<th>Variance</th>
<th>% R.E.</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Cum (\frac{3}{4}D_4(x, z)) Rule</td>
<td>Khan et al. (2008)</td>
</tr>
<tr>
<td>(1.7880, 1.5000)</td>
<td>0.0354952</td>
<td>233.64</td>
</tr>
<tr>
<td>(2.6870, 1.5000)</td>
<td>0.08293</td>
<td></td>
</tr>
<tr>
<td>(4.0000, 1.5000)</td>
<td>(1.7880, 2.0000)</td>
<td></td>
</tr>
<tr>
<td>(2.6870, 2.0000)</td>
<td>(4.0000, 2.0000)</td>
<td></td>
</tr>
</tbody>
</table>
12. Simulation Study

In this section, we conduct a simulation study to investigate the effectiveness of the proposed dynamic programming with the following methods (1-3) in stratification package in the R statistical software and 4 & 5 in LINGO:
1. Dalenius and Hodges [1959] cum f method, which is the most frequently used and better known method.
4. Khan et al. [2015] method
5. Proposed method.

In this study, a data set (with N = 5000) following the uniform and exponential distribution with a = 0, b=1, c= 0 and d = 2 was randomly generated by the R software. Then, the OSBs using the proposed method as discussed earlier are obtained for the three different number of strata, that is, (L,M) = (3,4). Then, the OSBs using the proposed method are obtained for (L,M) = (3,4). The OSBs are determined using cum f method, geometric method and the Lavallée-Hidiroglou (Kozak’s) method using the stratification package with CV = 0.75 and Khan et al. [10] and the proposed method using LINGO.

Table 5. The variance of variables for different stratification methods

<table>
<thead>
<tr>
<th>Stratification Method</th>
<th>Variance (in e-09)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Dalenius and Hodges [1959] cum f method</td>
<td>312.8371</td>
</tr>
<tr>
<td>Gunning and Horgan [2004] method</td>
<td>2891.916</td>
</tr>
<tr>
<td>Khan et al. [2015]</td>
<td>589.7021</td>
</tr>
<tr>
<td>Proposed method</td>
<td>203.107</td>
</tr>
</tbody>
</table>

From the table above, it is noted that the OSBs obtained by the cum f method and the proposed dynamic programming method are very close to each other, whereas the OBSs in the other methods, geometric, Lavallée-Hidiroglou method with Kozak’s algorithm and Khan et al. (2015) differ widely from that of the proposed method. However, the table reveals that the proposed method yields the smallest variances of the estimate for all (L, M) = (3,4) as compared to all the other methods. Although the variances for the dynamic programming method are closed to the cum f method, the other two methods produce a greater variance than the dynamic programming technique. Thus, the study reveals that the proposed dynamic programming technique
is more efficient than the other methods while stratifying a population with a uniform and exponential distributions.

13. Conclusion

The optimum stratification is defined as subdividing heterogeneous population into the best possible manner that makes the homogeneity within subpopulation and heterogeneity between them. Demarcation of strata boundaries is one of the main factors for efficient results in stratified random sampling. In this regard, we have proposed $\text{Cum } 3D(x,z)$ Rule ($i = 3, 4$) for obtaining approximately OSB for two stratification variables having single study variables for both the dependent as well as independent cases for concomitant variables. Thus, comparing the proposed method $\text{Cum } 3D(x,z)$ Rule for standard normally distributed auxiliary variables with the Singh (1975), the %RE obtained is 268.21, which indicates the efficiency of the proposed method. Further, the %RE obtained while making comparisons between the proposed method ($\text{Cum } 3D(x,z)$ Rule) and the method given by Yadava and Singh (1984) results in 169.89 for right-triangular and exponential auxiliary variables. In the same case for right-triangular and uniform auxiliary variables the %RE comes out to be 233.64 as compared with Khan et al. (2008) under proportional allocation. Further, the simulation study also proved the superiority of the proposed methods with regard to the existing methods. Thus, it can be concluded that the use of two stratification variables gains efficiency over a single auxiliary variable and the proposed methods are more precise than the existing methods. The proposed strategy can be entirely applied to different distributions that describe the concomitant variables.

References


